

Problem: Inference in generative models

Given: Probabilistic model of

\mathbf{y} : observed variables $\in \mathcal{Y} = \mathbb{R}^{N_y}$,

\mathbf{z} : latent variables $\in \mathcal{Z} = \mathbb{R}^{N_z}$,

where we can only generate (\mathbf{y}, \mathbf{z}) pairs.

Task: Estimate conditional expectations

$$\mathbb{E}[f(\mathbf{z}) | \mathbf{y} = \mathbf{y}_{\text{obs}}] =$$

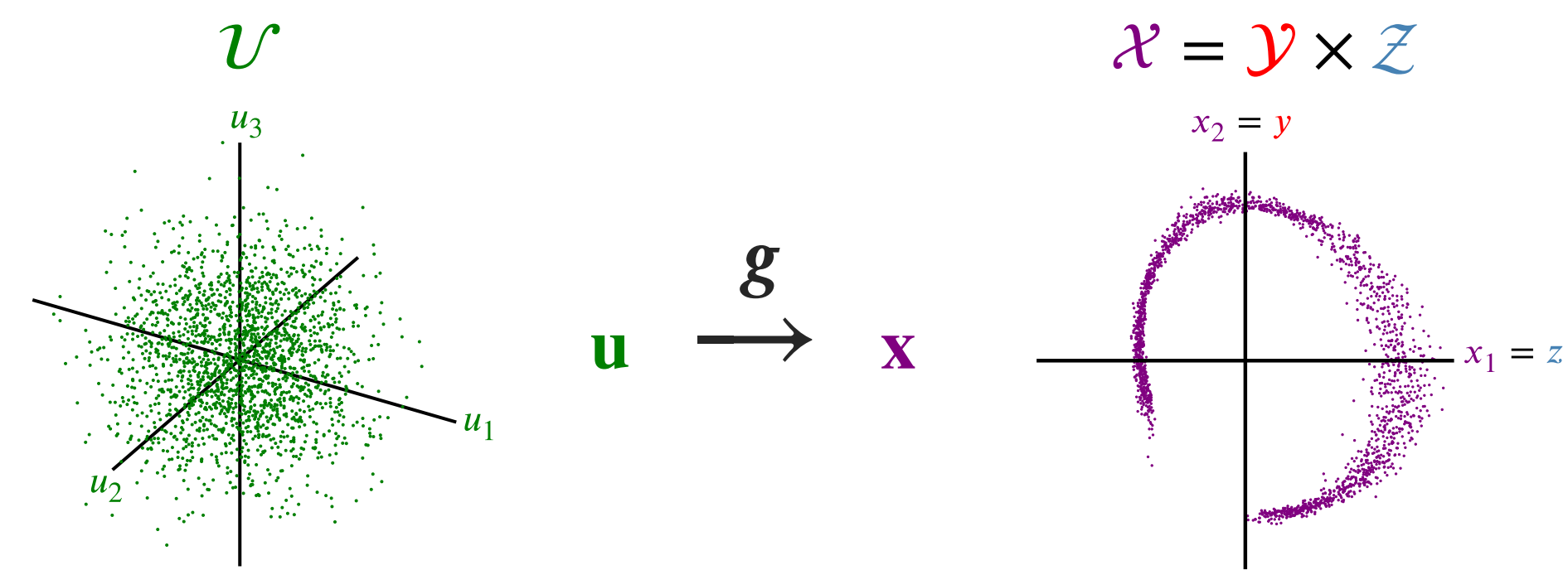
$$\int_{\mathcal{Z}} f(\mathbf{z}) \mathbb{P}_{\mathbf{z} | \mathbf{y}}[\mathbf{z} | \mathbf{y}_{\text{obs}}] d\mathbf{z},$$

of latent variables given observations \mathbf{y}_{obs} .

Differentiable generative models

Generative model, a.k.a. implicit or simulator model:

A model where we can independently sample (\mathbf{y}, \mathbf{z}) but not necessarily evaluate $\mathbb{P}_{\mathbf{y}, \mathbf{z}}$.



A generative model can be expressed in the form $\mathbf{u} \sim \rho$, $\mathbf{x} = \mathbf{g}(\mathbf{u})$ with

\mathbf{u} : random inputs $\in \mathcal{U}$,

$\rho: \mathcal{U} \rightarrow \mathbb{R}^+$: input density,

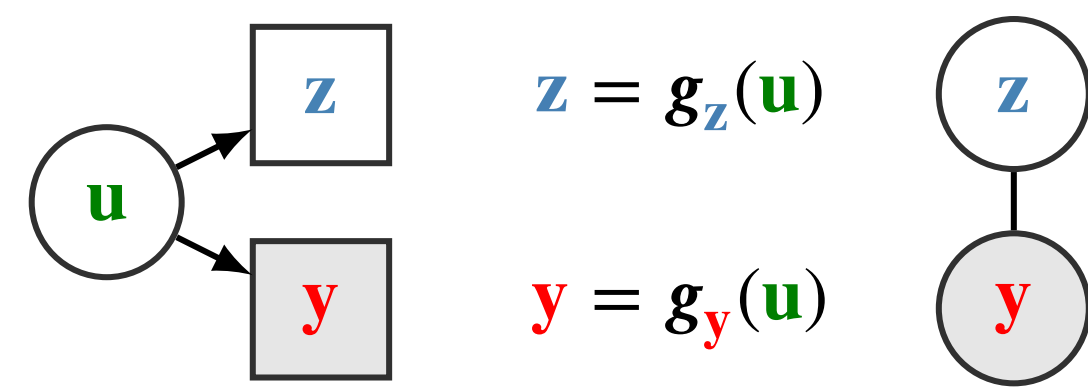
$\mathbf{x} = (\mathbf{y}, \mathbf{z})$: generated outputs $\in \mathcal{X} = \mathcal{Y} \times \mathcal{Z}$,

$\mathbf{g}: \mathcal{U} \rightarrow \mathcal{X}$: generator function.

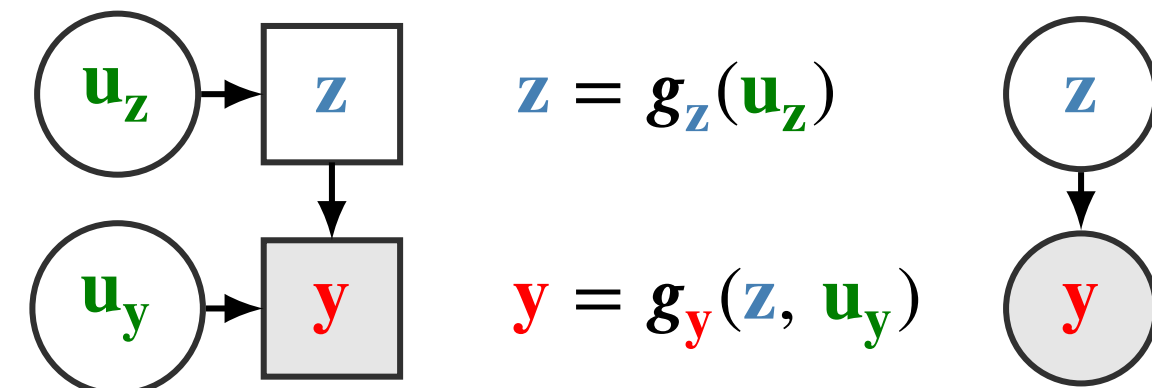
Here we will consider models where $\mathcal{U} = \mathbb{R}^M$ and \mathbf{g} is differentiable i.e. $\frac{\partial \mathbf{g}}{\partial \mathbf{u}}$ exists a.e.

Undirected and directed models

Undirected generative model

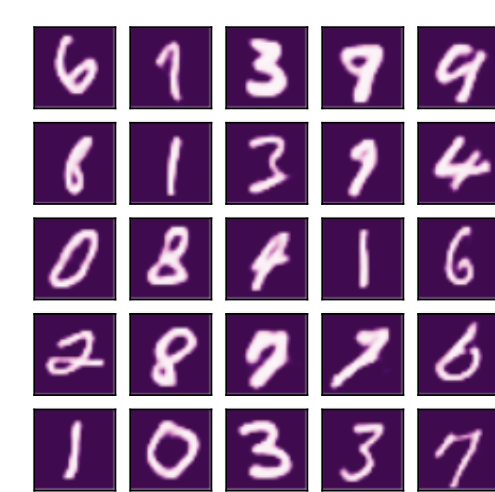


Directed generative model



Examples of differentiable generative models

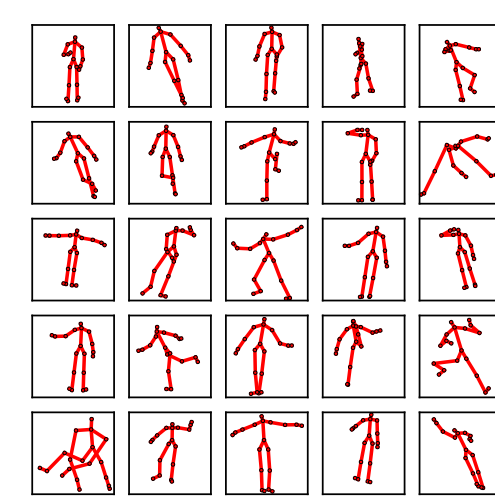
In all the examples below the input density is Gaussian $\rho(\mathbf{u}) = \mathcal{N}(\mathbf{u}; \mathbf{0}, \mathbf{I})$.



Decoder of a Gaussian Variational Autoencoder

$$\mathbf{x} = \mathbf{m}(\mathbf{u}_1) + \mathbf{s}(\mathbf{u}_1) \odot \mathbf{u}_2,$$

with \mathbf{m} and \mathbf{s} parametric functions trained to match $\mathbb{P}_{\mathbf{x}}$ to the distribution of a dataset e.g. MNIST digit images.

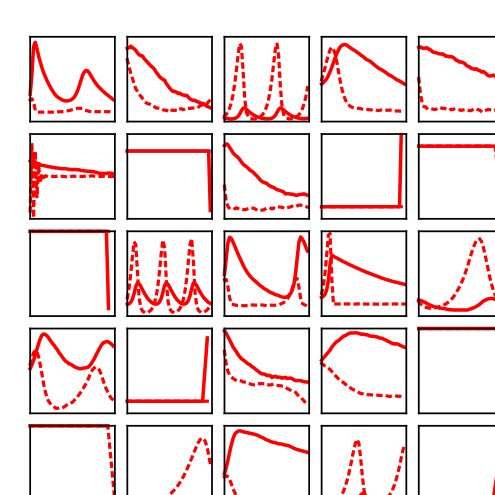


Pose generator, with learnt 3D pose parameter generators

$$\mathbf{z}_a = \mathbf{f}_a(\mathbf{u}_a) \quad \mathbf{z}_b = \mathbf{f}_b(\mathbf{u}_b) \quad \mathbf{z}_c = \mathbf{f}_c(\mathbf{u}_c)$$

and a pin-hole camera model to generate observed 2D projections

$$2\text{D proj. } \mathbf{y}_j = \mathbf{C}(\mathbf{z}_c) \mathbf{r}_j(\mathbf{z}_a, \mathbf{z}_b) + \sigma \mathbf{u}_j \quad \forall j \in \{1 \dots J\}.$$



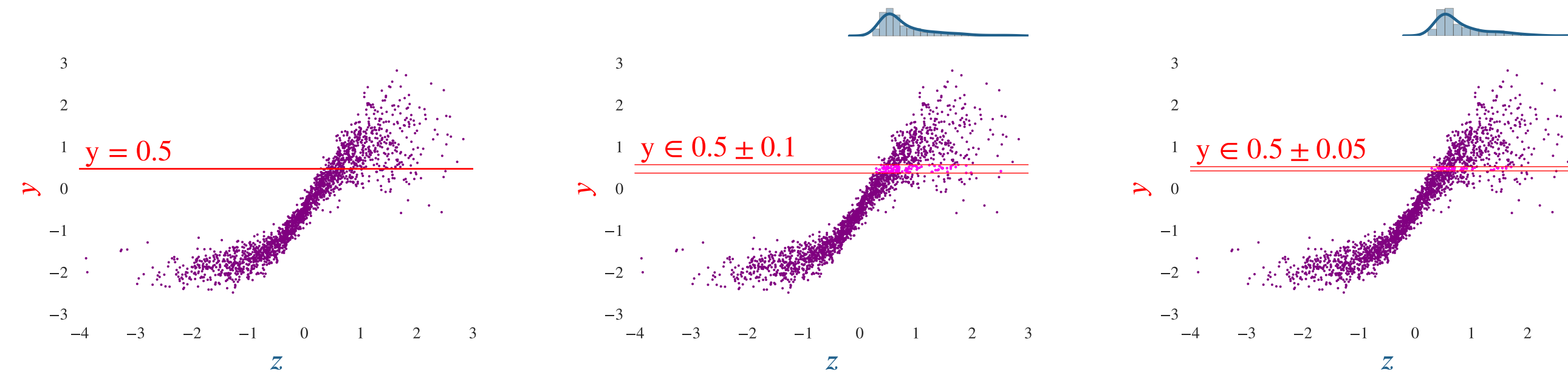
Continuous Lotka-Volterra predator-prey population model,

$$\begin{aligned} d\mathbf{y}_1 &= (z_1 \mathbf{y}_1 - z_2 \mathbf{y}_1 \mathbf{y}_2) dt + d\mathbf{n}_1, \\ d\mathbf{y}_2 &= (-z_3 \mathbf{y}_2 + z_4 \mathbf{y}_1 \mathbf{y}_2) dt + d\mathbf{n}_2. \end{aligned}$$

Simulator for system can be expressed as a directed model

$$\begin{aligned} \mathbf{z} &= \mathbf{g}_z(\mathbf{u}_z) = \exp(\sigma \odot \mathbf{u}_z - \mu): \text{ log-normal prior,} \\ \mathbf{y} &= \mathbf{g}_y(\mathbf{z}, \mathbf{u}_y): \text{ Euler-Maruyama integration of SDEs.} \end{aligned}$$

Approximate Bayesian Computation (ABC)



Family of approximate inference methods for generative models.

Typically applied to directed models where likelihood $\mathbb{P}_{\mathbf{y} | \mathbf{z}}$ is unavailable.

Key idea: true observations $\bar{\mathbf{y}}$ are decoupled from simulated observed values \mathbf{y} by a noise kernel $\mathbb{P}_{\bar{\mathbf{y}} | \mathbf{y}}[\bar{\mathbf{y}} | \mathbf{y}] = k_\epsilon(\bar{\mathbf{y}}; \mathbf{y})$. Common choices for the kernel include

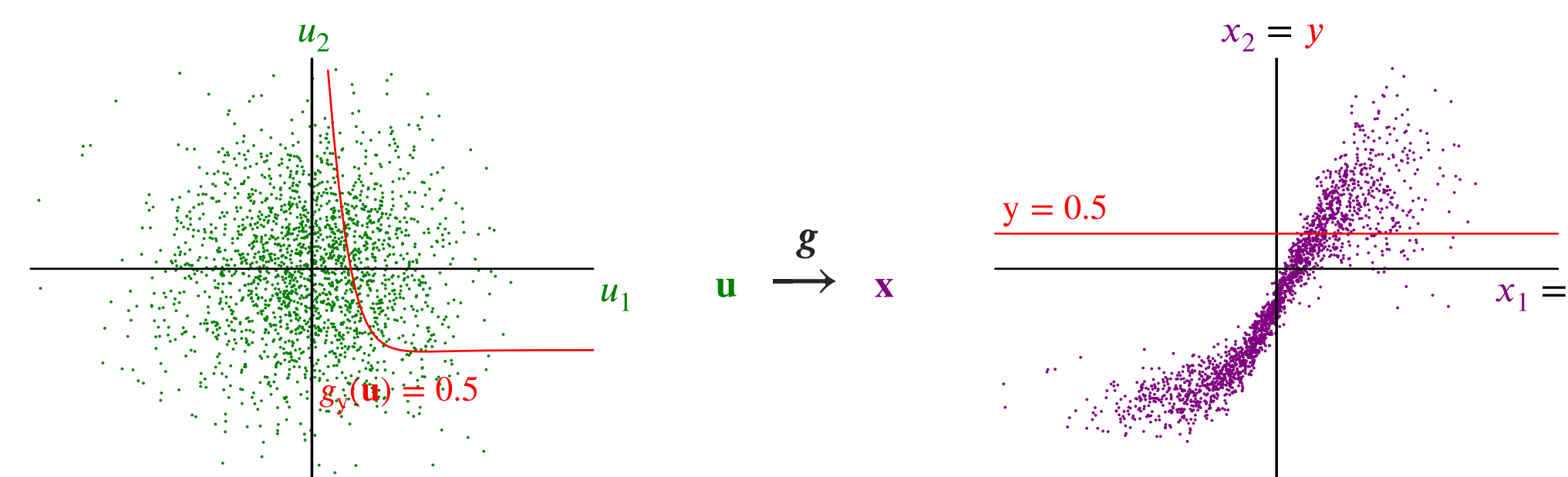
$$k_\epsilon(\bar{\mathbf{y}}; \mathbf{y}) \propto \mathbb{1}[\|\bar{\mathbf{y}} - \mathbf{y}\| < \epsilon] \quad (\text{uniform ball}), \quad k_\epsilon(\bar{\mathbf{y}}; \mathbf{y}) = \mathcal{N}(\bar{\mathbf{y}}; \mathbf{y}, \epsilon^2 \mathbf{I}) \quad (\text{Gaussian}).$$

Kernel can be used to express approximate conditional expectations

$$\mathbb{E}[f(\mathbf{z}) | \bar{\mathbf{y}} = \mathbf{y}_{\text{obs}}; \epsilon] = \frac{1}{C} \iint_{\mathcal{Y} \times \mathcal{Z}} f(\mathbf{z}) k_\epsilon(\mathbf{y}_{\text{obs}}; \mathbf{y}) \mathbb{P}_{\mathbf{y}, \mathbf{z}}[\mathbf{y}, \mathbf{z}] d\mathbf{y} d\mathbf{z},$$

with asymptotic consistence in that $\lim_{\epsilon \rightarrow 0} \mathbb{E}[f(\mathbf{z}) | \bar{\mathbf{y}} = \mathbf{y}_{\text{obs}}; \epsilon] = \mathbb{E}[f(\mathbf{z}) | \mathbf{y} = \mathbf{y}_{\text{obs}}]$.

Inference in the generator input space



The Law of the Unconscious Statistician can be used to rewrite the ABC conditional expectation as an integral over the generator input space \mathcal{U}

$$\mathbb{E}[f(\mathbf{z}) | \bar{\mathbf{y}} = \mathbf{y}_{\text{obs}}; \epsilon] = \frac{1}{C} \int_{\mathcal{U}} f \circ \mathbf{g}_z(\mathbf{u}) k_\epsilon(\mathbf{y}_{\text{obs}}; \mathbf{g}_y(\mathbf{u})) \rho(\mathbf{u}) d\mathbf{u}.$$

By taking the $\epsilon \rightarrow 0$ limit of the above and applying Federer's Co-Area Formula the exact conditional expectation can be expressed as an integral

$$\mathbb{E}[f(\mathbf{z}) | \mathbf{y} = \mathbf{y}_{\text{obs}}] = \frac{1}{C} \int_{\mathcal{M}_{\mathbf{y}_{\text{obs}}}} f \circ \mathbf{g}_z(\mathbf{u}) \left| \frac{\partial \mathbf{g}_y}{\partial \mathbf{u}} \frac{\partial \mathbf{g}_y}{\partial \mathbf{u}}^T \right|^{-\frac{1}{2}} \rho(\mathbf{u}) \mathcal{H}_{\mathcal{M}_{\mathbf{y}_{\text{obs}}}} \{d\mathbf{u}\}.$$

over an implicitly defined manifold $\mathcal{M}_{\mathbf{y}_{\text{obs}}} = \{\mathbf{u} \in \mathcal{U} : \mathbf{g}_y(\mathbf{u}) = \mathbf{y}_{\text{obs}}\}$ embedded in \mathcal{U} .

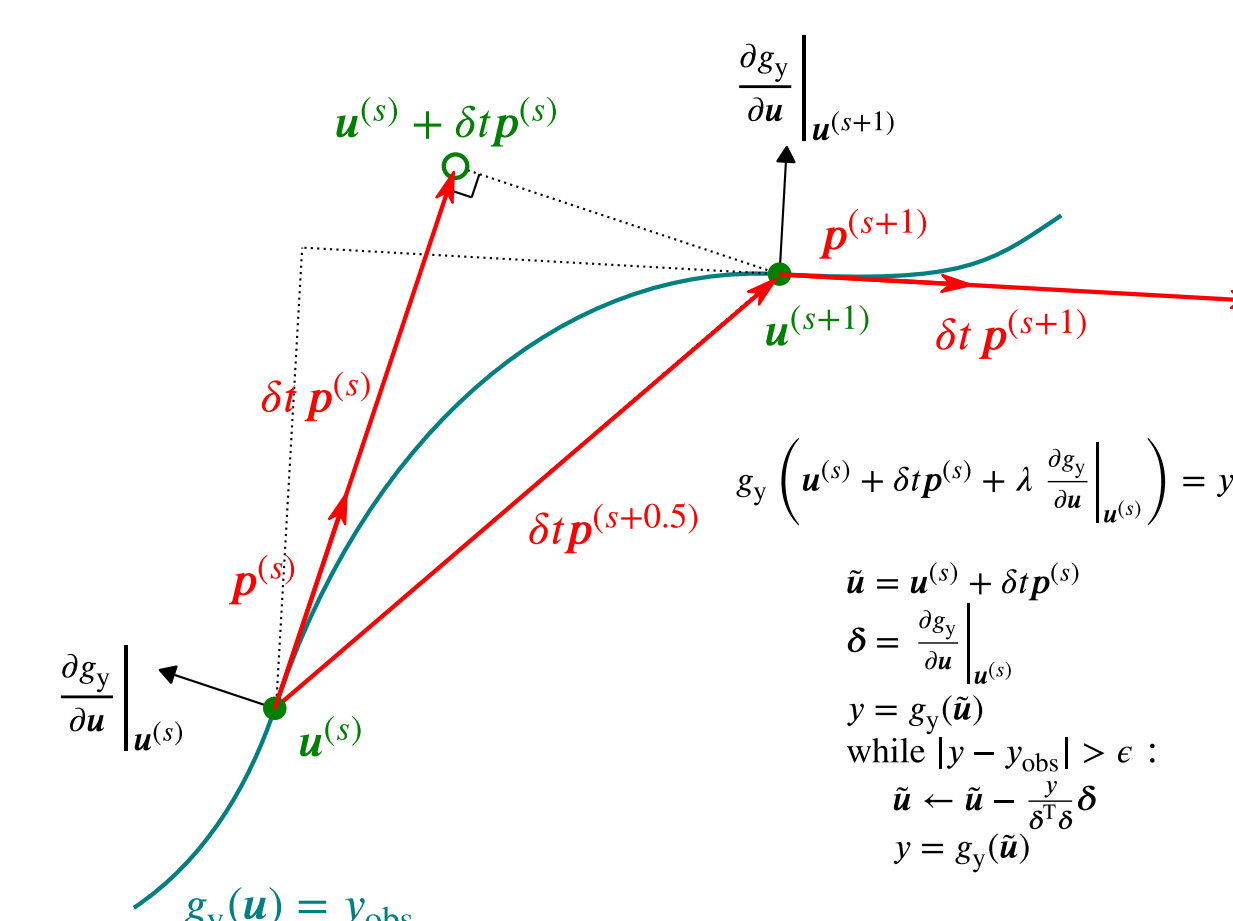
Constrained Hamiltonian Monte Carlo (CHMC)

Use simulated constrained Hamiltonian dynamic to propose moves on $\mathcal{M}_{\mathbf{y}_{\text{obs}}}$:

$$\pi(\mathbf{u}) = \frac{1}{C} \left| \frac{\partial \mathbf{g}_y}{\partial \mathbf{u}} \frac{\partial \mathbf{g}_y}{\partial \mathbf{u}}^T \right|^{-\frac{1}{2}} \rho(\mathbf{u})$$

$$\frac{d\mathbf{u}}{dt} = \mathbf{p} \quad \frac{d\mathbf{p}}{dt} = \frac{\partial \log \pi}{\partial \mathbf{u}} - \frac{\partial \mathbf{g}_y}{\partial \mathbf{u}}^T \lambda$$

subject to $\mathbf{g}_y(\mathbf{u}) = \mathbf{y}_{\text{obs}}$ and $\frac{\partial \mathbf{g}_y}{\partial \mathbf{u}} \mathbf{p} = \mathbf{0}$.

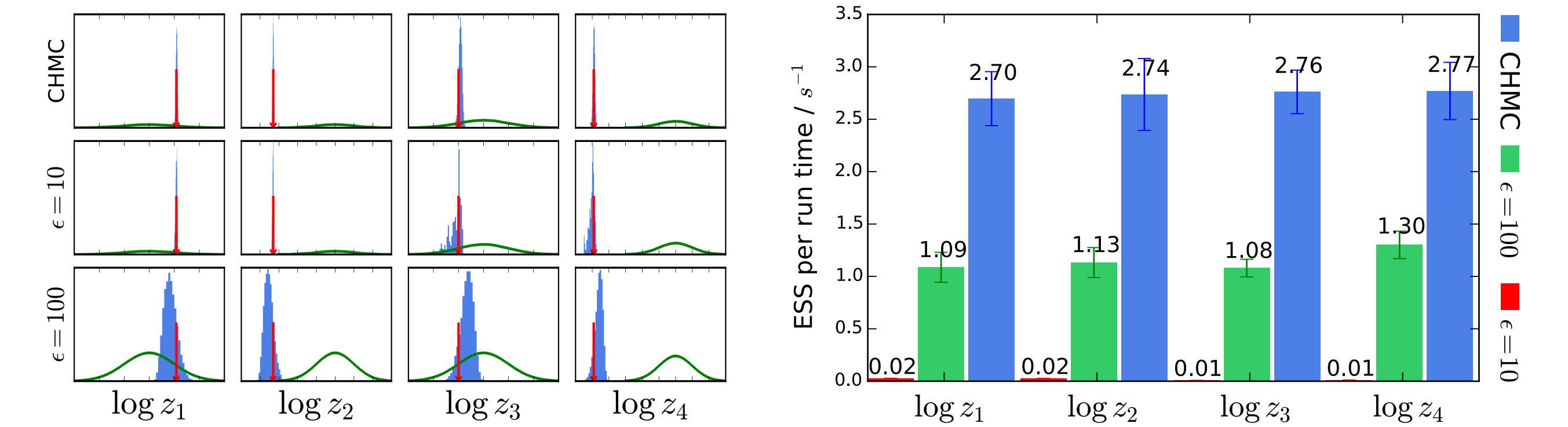
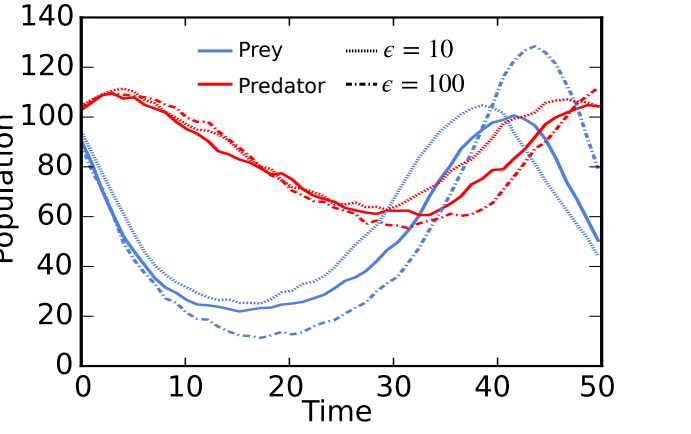


Experiments

Lotka-Volterra parameter inference

Infer posterior on 4 system parameters given observations of predator and prey populations at 50 simulated time-steps.

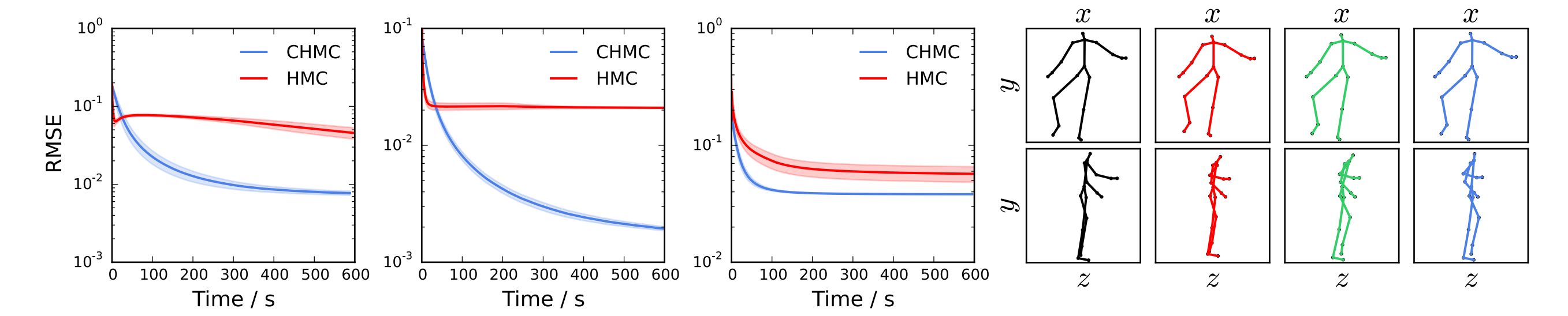
Pseudo-marginal slice sampling in ABC posterior with uniform ball kernel with $\epsilon = 10$ and $\epsilon = 100$ used as baseline.



Human pose inference

3D pose given binocular 2D projections

3D pose given monocular 2D projection



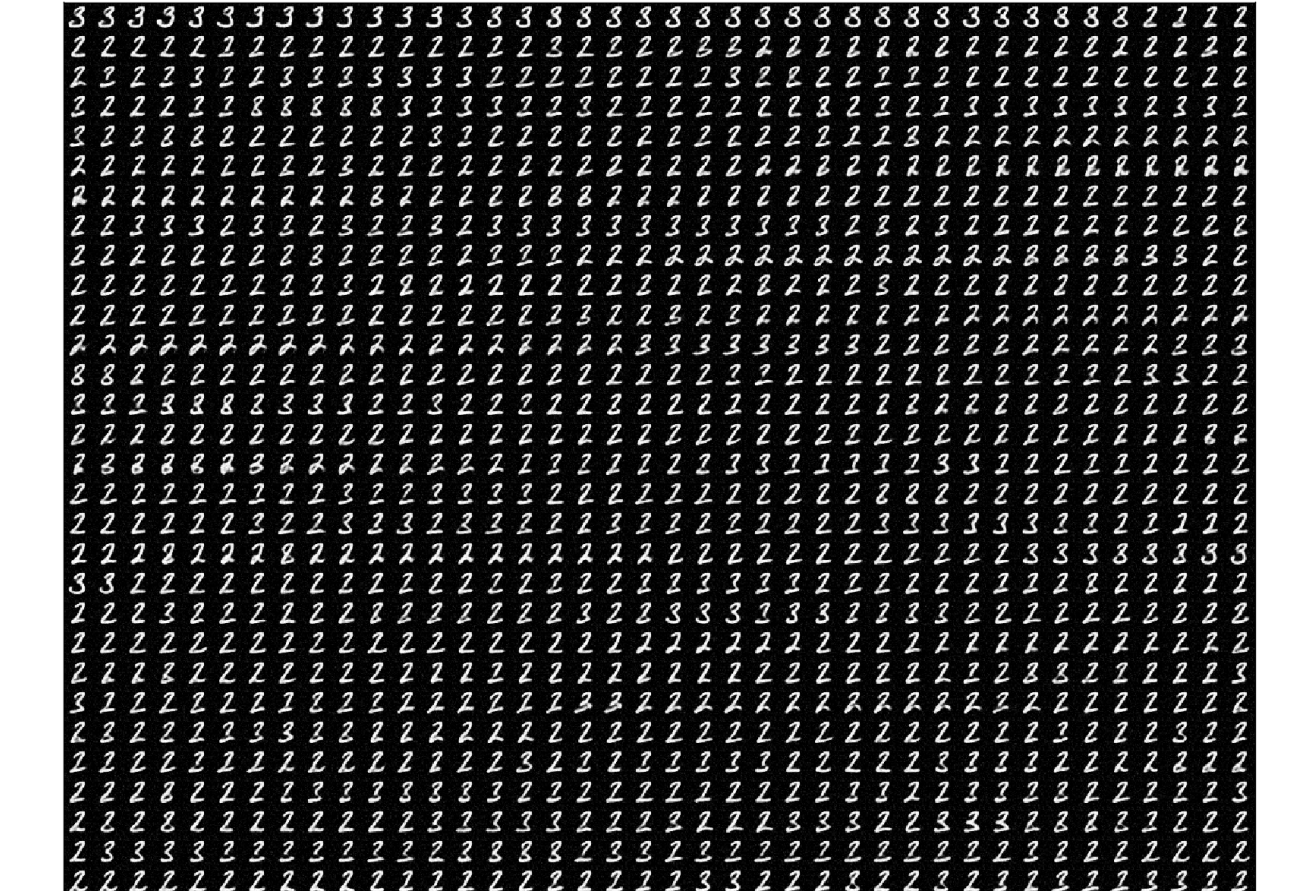
MNIST digit image in-painting

In-painting bottom 75% of digit images given observed top 25%. 60 s total sampling time.

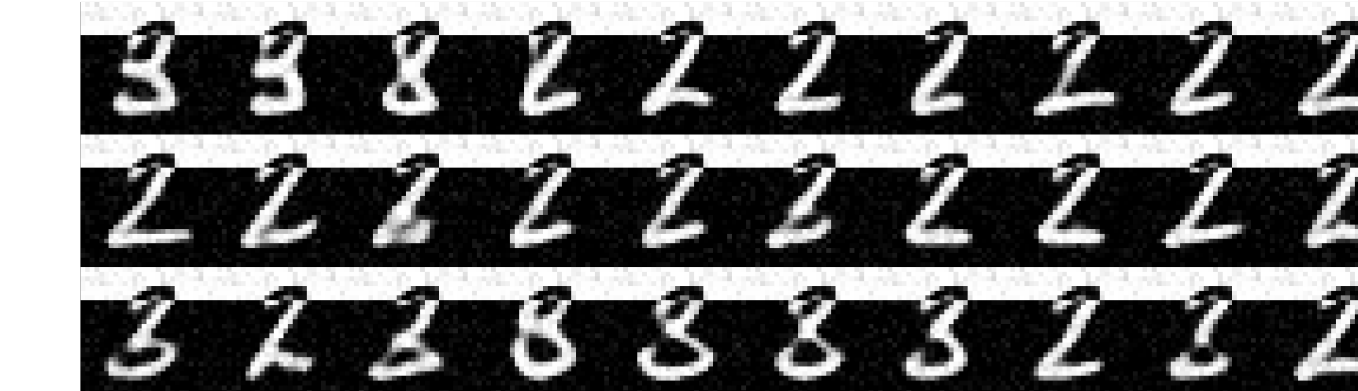
Consecutive constrained HMC in-painted image samples



Consecutive standard HMC in-painted image samples



Factor 40 thinned standard HMC in-painted image samples



Conclusions

Generally applicable inference method for differentiable generative models.

Asymptotically exact alternative to ABC where applicable:
 $\epsilon = 0$ / no summary statistics.

Key idea: consider conditioning as constraint on inputs to generator function.

Constrained HMC allow efficient gradient-based exploration of target density on constraint manifold corresponding to input consistent with observations.

References

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